

# Estimating the parametric covariation matrix: Equivalence, efficiency and estimation

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# Introduction

Consider the  $d$ -dimensional observation model

$$Y_i = X_{i/n} + \varepsilon_i, \quad i = 1, \dots, n,$$

where  $X_t = \Sigma^{1/2} G_t$  with  $G^{(j)} \stackrel{i.i.d.}{\sim} \Gamma$ ,  $j = 1, \dots, d$ , for  $\Gamma$  Gaussian.

The target of estimation is the quadratic covariation matrix  $\Sigma \in \mathbb{R}^{d \times d}$  and belongs to the class

$$\mathfrak{S}_L = \{\Sigma \in \mathbb{R}_+^{d \times d} : \Sigma_0 \leq \Sigma, \|\Sigma\| \leq L\}.$$

The errors  $\varepsilon_i$  are i.i.d.  $\mathcal{N}(0, \eta^2 I_d)$  distributed and independent of  $G$ .

# Possible application: High-frequency data

Literature concerning microstructure noise includes works by Aït-Sahalia, Andersen, Barndorff-Nielsen, Bibinger, Christensen, Fan, Gloter, Hansen, Hautsch, Hoffmann, Jacod, Li, Lunde, Mykland, Munk, Podolskij, Reiß, Rosenbaum, Schmidt-Hieber, Shephard, Todorov, Uchida, Vetter, Yoshida, Zhang, Zheng.

## Many estimation approaches

realised covariances, quasi Maximum likelihood, realised kernels, preaveraging, scaling, spectral estimators, etc.

## What about explicit lower bounds?

$d = 1$ :

Case  $\Gamma = B$

- Gloter and Jacod [2001]:  $n^{1/4}(\hat{\sigma}^2 - \sigma^2) \xrightarrow{d} \mathcal{N}(0, 8\eta\sigma^3)$

Case  $\Gamma = B^H$

- Gloter and Hoffmann [2007]:  $r_n = n^{-1/(4H+2)}$
- Sabel and Schmidt-Hieber [2014]: Cramér-Rao bound

$d > 1$ :

Case  $\Gamma = B$

- Bibinger et al. [2014]: Cramér-Rao bound

Here: General approach for a wide class of  $\Gamma$ .

## Interlude: Le Cam theory

### Definition

Let  $\mathcal{E} = (X, \mathcal{X}, \{P_\theta : \theta \in \Theta\})$  and  $\mathcal{F} = (Y, \mathcal{Y}, \{Q_\theta : \theta \in \Theta\})$  be two statistical experiments. Then the Le Cam deficiency between  $\mathcal{E}$  and  $\mathcal{F}$  is given by

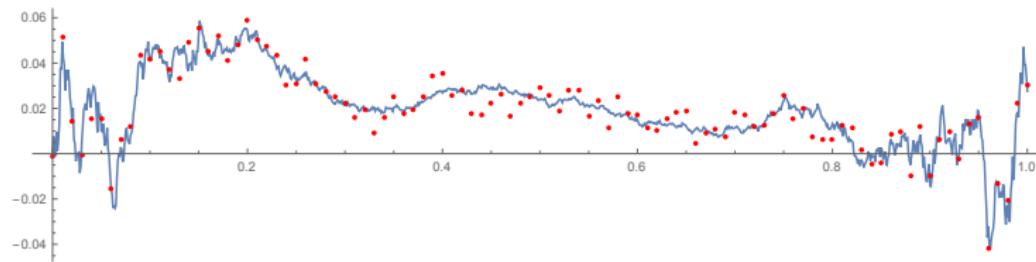
$$\delta(\mathcal{E}, \mathcal{F}) = \inf_K \sup_{\theta \in \Theta} \|K \cdot P_\theta - Q_\theta\|_{TV},$$

where the infimum is taken over all Markov kernels from  $(X, \mathcal{X})$  to  $(Y, \mathcal{Y})$ .

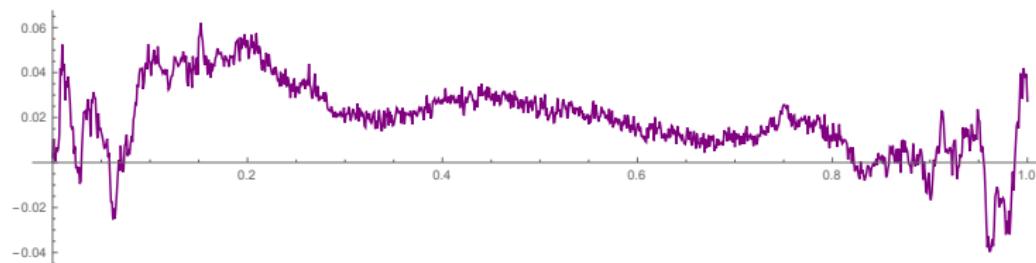
The Le Cam distance is defined by  $\Delta(\mathcal{E}, \mathcal{F}) = \max\{\delta(\mathcal{E}, \mathcal{F}), \delta(\mathcal{F}, \mathcal{E})\}$ .

- If  $\delta(\mathcal{E}, \mathcal{F}) = 0$  we say that  $\mathcal{E}$  is more informative than  $\mathcal{F}$ .
- If  $\Delta(\mathcal{E}, \mathcal{F}) = 0$  we say that  $\mathcal{E}$  and  $\mathcal{F}$  are equivalent.

## Discrete and continuous model



$$\mathcal{D}_n : Y_i = X_{i/n} + \varepsilon_i, \quad i = 1, \dots, n.$$



$$\mathcal{C}_n : dY_t^n = X_t dt + \eta n^{-1/2} dW_t, \quad t \in [0, 1]$$

## Proposition

Denote by  $c(s, t)$  the covariance function of  $\Gamma$ . Assume that the variance function  $v(t) = c(t, t)$  belongs to  $C^\gamma([0, 1])$  for  $\gamma > 1/2$ . Additionally, suppose that for every  $s \in [0, 1]$  the derivative of  $c_s(t) = c(s, t)$  lies in  $C^\beta([0, s)) \cap C^\beta((s, 1])$ . Then it holds

$$\Delta(\mathcal{D}_n, \mathcal{C}_n) = \mathcal{O}(n^{-(\beta \wedge 1/2 \wedge (\gamma - 1/2))}).$$

## Corollary

Let  $\Gamma = B^H$  be a fractional Brownian motion with Hurst exponent  $H > 1/2$ . Then the corresponding discrete and continuous experiment are asymptotically equivalent. More precisely,

$$\Delta(\mathcal{D}_n, \mathcal{C}_n) = \mathcal{O}(n^{-((2H-1) \wedge 1/2)}).$$

# Sequence space model

## Some notations

- $C_\Gamma$ : covariance operator of  $\Gamma$ , given by

$$f(t) \mapsto \int_0^1 c(s, t) f(s) ds$$

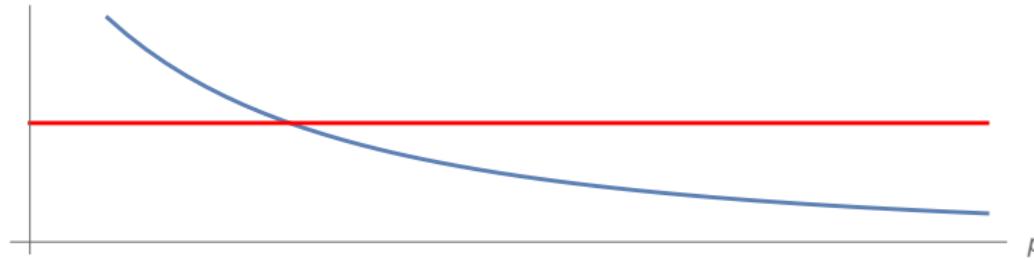
- denote by  $\lambda = (\lambda_p)_{p \geq 1}$  and  $e_p$  the eigenvalues and eigenfunctions of  $C_\Gamma$ , respectively

**Now:** Restrict  $\mathcal{C}_n$ :  $dY_t^n = X_t dt + n^{-1/2} dW_t$ , to the eigenfunctions  $e_p$  and obtain a representation in terms of an independent sequence

$$Y_{np} \sim \mathcal{N}(0, \Sigma \lambda_p + \frac{1}{n} I_d), \quad p \geq 1. \quad (\mathcal{S}_n)$$

# Equilibrium

Consider the variances  $\Sigma \lambda_p + \frac{1}{n} I_d$  and the following plot of  $\lambda$  and  $\frac{1}{n}$ :



Equilibrium point:  $\lambda(P_n) = \frac{1}{n}$ :

## Assumption

$\lambda$  is regularly varying at infinity with index  $-\alpha$ ,  $\alpha > 1$ .

## Theorem

For  $\lambda(P_n) = \frac{1}{n}$  the Fisher information satisfies

$$P_n^{-1} \sum_{p=1}^{\infty} \mathcal{I}_{np}(\Sigma) \rightarrow \mathcal{I}(\Sigma),$$

uniformly in  $\Sigma \in \mathfrak{S}_L$ . The asymptotic Fisher information matrix  $\mathcal{I}(\Sigma)$  equals

$$\frac{\pi C}{4\alpha \sin(\pi/\alpha)} (Q^{\otimes 2})^\top \text{diag} \left\{ \text{diag} \left\{ \frac{s_j^{1/\alpha-1} - s_i^{1/\alpha-1}}{s_i - s_j} \right\}_{1 \leq j \leq d} \right\}_{1 \leq i \leq d} Q^{\otimes 2},$$

with  $Q \in \mathbb{R}^{d \times d}$  being the orthogonal matrix such that  $\Sigma = Q^\top \text{diag}\{s_i\}_{1 \leq i \leq d} Q$  and  $C = \lim_{n \rightarrow \infty} r_n^2 P_n$ .

# Consequences

- the optimal rate for estimating  $\Sigma$  efficiently is given by

$$r_n \sim P_n^{-1/2}$$

- in particular, for some slowly varying  $L$

$$r_n = \mathcal{O}(n^{-1/(2\alpha)} L(n)^{-1/2})$$

- asymptotically sufficient information for estimating  $\Sigma$  efficiently is contained in intervals  $(P_n \kappa_n^0, P_n \kappa_n^\infty)$  for  $\kappa_n^0 \rightarrow 0$ ,  $\kappa_n^\infty \rightarrow \infty$
- $\alpha$  as smoothness index: smoothness  $\uparrow \Rightarrow$  rate  $\downarrow$   
(It even holds:  $\lambda > \lambda' \Rightarrow \delta(\mathcal{S}_n, \mathcal{S}'_n) = 0$ )

## Theorem

The model  $\mathcal{S}_n$  possesses the LAN property, i.e.

$$\begin{aligned} \log \frac{dP_{\Sigma+r_n H, n}}{dP_{\Sigma, n}}(Y_n) = & r_n \text{vec}(H)^\top \nabla \ell(Y_n, \Sigma) \\ & - r_n^2 \frac{1}{2} \text{vec}(H)^\top \mathcal{I}_n(\Sigma) \mathcal{Z} \text{vec}(H) + \rho_n, \end{aligned}$$

where  $r_n \nabla \ell(Y_n, \Sigma) \xrightarrow{d} \mathcal{N}(0, \mathcal{I}(\Sigma) \mathcal{Z})$  and  $\rho_n = o_p(1)$ .

## Corollary

A sequence  $\text{vec}(\hat{\Sigma}_n)$  of efficient regular estimators of  $\text{vec}(\Sigma)$  satisfies

$$r_n^{-1} \text{vec}(\hat{\Sigma}_n - (\Sigma + r_n H)) \rightarrow \mathcal{N}(0, \frac{1}{4} \mathcal{I}(\Sigma)^{-1} \mathcal{Z}),$$

under  $P_{\Sigma+r_n H}$  and for a certain normalisation matrix  $\mathcal{Z} \in \mathbb{R}^{d^2 \times d^2}$ .

# Common lower bounds

Consider the underlying eigenvalues of the Brownian bridge and the Brownian motion:

$$\lambda_p^{BB} = (\pi p)^{-2} \quad \text{and} \quad \lambda_p^{BM} = \pi^{-2}(p + 1/2)^{-2}.$$

**Question:** Do the lower bounds coincide?

## Proposition

Let  $\lambda, \lambda' \in RV(-\alpha)$  induce the models  $\mathcal{S}_n$  and  $\mathcal{S}'_n$ , respectively.  
Assume  $|\lambda_p - \lambda'_p| = \mathcal{O}(p^{-(\alpha+\varepsilon)})$  with  $\varepsilon > 1/2$ . Then

$$\Delta(\mathcal{S}_n, \mathcal{S}'_n) = \mathcal{O}\left(P_n^{1/2-\varepsilon}\right).$$

What if  $\varepsilon \leq 1/2$ ?

# Common lower bounds II

## Lemma

Let  $\mathcal{S}_n$  and  $\mathcal{S}'_n$  be two models induced by  $\lambda, \lambda'$ , respectively. Then

$$\lambda_p/\lambda'_p \rightarrow 1 \iff r_n \sim r'_n \text{ and } \mathcal{I}(\Sigma) = \mathcal{I}'(\Sigma).$$

In particular, it holds

$$r_n^2 \sum_{p=1}^{\infty} \|\mathcal{I}_{np}(\Sigma) - \mathcal{I}'_{np}(\Sigma)\| = \mathcal{O}\left(\frac{\lambda_{P_n} - \lambda'_{P_n}}{\lambda_{P_n}}\right).$$

## Example

Consider the Brownian motion, the Brownian bridge and the Ornstein-Uhlenbeck process. The underlying eigenvalues are

$$\lambda_p^{BM} = \pi^{-2}(p - 1/2)^{-2}, \quad \lambda_p^{BB} = (\pi p)^{-2}, \quad \lambda_p^{OU} = (\pi^2 p^2 + \theta^2)^{-1}.$$

Thus their lower bounds coincide and the corresponding central limit theorem is given by

$$n^{1/4} \text{vec}(\hat{\Sigma}_n - \Sigma) \rightarrow \mathcal{N}(0, 2(\Sigma \otimes \Sigma^{1/2} + \Sigma^{1/2} \otimes \Sigma)\mathcal{Z}).$$

## Example

Consider the case  $\Gamma = B^H$ ,  $0 < H < 1$ . Due to Bronski [2003] and Chigansky and Kleptsyna [2016] the leading terms of the eigenvalues of  $C_\Gamma$  are known:

$$\lambda_p^{fBM} = \frac{\sin(\pi H)\Gamma(2H+1)}{(p\pi)^{2H+1}} + o(p^{-(2H+1)}).$$

For  $d = 1$  the corresponding central limit theorem is given by

$$n^{\frac{1}{(4H+2)}} (\hat{\sigma}_n^2 - \sigma^2) \rightarrow \mathcal{N}\left(0, \sigma^{\frac{2(4H+1)}{(2H+1)}} \cdot \frac{(2H+1)^2 \cdot \sin(\pi/(2H+1))}{H(\sin(H\pi) \cdot \Gamma(2H+1))^{1/(2H+1)}}\right).$$

Others: fOU, (fractional) Gaussian sheets, etc.

# Oracle estimator

For every observation  $Y_{np}$  construct an estimator of  $\text{vec}(\Sigma)$  via

$$\text{vec}(\hat{\Sigma}_{np}) = \lambda_p^{-1} \text{vec} \left( Y_{np} Y_{np}^\top - \frac{1}{n} I_d \right).$$

With an optimal choice of (oracle) weights  $W_{np}^{Or}(\Sigma)$  set

$$\text{vec}(\hat{\Sigma}_n^{Or}) = \sum_{p=1}^{P_n S_n} W_{np}^{Or}(\Sigma) \text{vec}(\hat{\Sigma}_{np}).$$

## Theorem

If  $S_n \rightarrow \infty$  then, under  $P_{\Sigma + r_n H, n}$ , the oracle estimator satisfies

$$r_n^{-1} \text{vec}(\hat{\Sigma}_n^{Or} - (\Sigma + r_n H)) \rightarrow \mathcal{N}(0, \frac{1}{4} \mathcal{I}(\Sigma)^{-1} \mathcal{Z}).$$

# Adaptive estimator

Assume that a consistent pre-estimate  $\text{vec}(\hat{\Sigma}_n^{pre})$  of  $\text{vec}(\Sigma)$  can be derived. Plug-in:

$$\text{vec}(\hat{\Sigma}_n^{Ad}) = \sum_{p=1}^{P_n S_n} W_{np}^{Or}(\hat{\Sigma}_n^{pre}) \text{vec}(\hat{\Sigma}_{np}).$$

## Theorem

Let a pre-estimator  $\text{vec}(\hat{\Sigma}_n)$  with  $\|\hat{\Sigma}_n - \Sigma\| = o_p(1)$ . Then under  $P_{\Sigma + r_n H, n}$

$$r_n^{-1} \text{vec}(\hat{\Sigma}_n^{Ad} - (\Sigma + r_n H)) \rightarrow \mathcal{N}(0, \frac{1}{4} \mathcal{I}(\Sigma)^{-1} \mathcal{Z}).$$

## Proposition

The MLE is a consistent estimator.

# Outlook

- extend the class of model (drifts, time-dependent  $\Sigma$ , etc.)
- find the boundaries of the Le Cam result ( $\Gamma = B^H$ ,  $H = 1/4?$ )
- sequence space approximations if  $\lambda$  not completely known

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