

VARIANCE FUNCTION OF WISHART EXPONENTIAL FAMILIES IN GRAPHICAL MODELS.

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MOTIVATION

- Graphical model: a statistical model + graph representation.
Variable = Vertex. No edge = conditional independence.
- Example: $A_n : 1 - 2 - 3 - \dots - n$.
For all $|i - j| > 1$, $X_i \perp X_j | (X_k)_{k \neq i, j}$.
- In the case of a Gaussian graphical model with covariance matrix Σ and concentration matrix $K = \Sigma^{-1}$, the conditional independence constraints are equivalent to $K_{ij} = 0$ for all non adjacent i and j .

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- This implies that the elements Σ_{ij} are not free parameters. The model can be parametrised by an incomplete symmetric matrix $\pi(\Sigma)$ with entries corresponding to non-adjacent vertices left out and such that the submatrices corresponding to complete subsets of vertices are positive definite. This set of incomplete matrices is denoted Q_G .

EXAMPLE

For the graph $1 - 2 - 3 - 4$:

The elements of Q_{A_4} are incomplete matrices of the form

$$x = \begin{pmatrix} x_{11} & x_{12} & * & * \\ x_{12} & x_{22} & x_{23} & * \\ * & x_{23} & x_{33} & x_{34} \\ * & * & x_{34} & x_{44} \end{pmatrix}$$

such that $\begin{vmatrix} x_{11} & x_{12} \\ x_{12} & x_{22} \end{vmatrix}$, $\begin{vmatrix} x_{22} & x_{23} \\ x_{23} & x_{33} \end{vmatrix}$ and $\begin{vmatrix} x_{33} & x_{34} \\ x_{34} & x_{44} \end{vmatrix}$ are positive definite.

OBJECTIVE

Derive the variance function the exponential family of Wishart distributions constructed on the cone Q_{A_n} ,

- The mean function is given by

$$m_{\underline{s}}^{(M)}(y) = \pi \left(\sum_{i=1}^{M-1} (s_i - s_{i+1}) [(y_{\{1:i\}})^{-1}]^0 + s_M y^{-1} + \sum_{i=M+1}^n (s_{i+1} - s_i) [(y_{\{i:n\}})^{-1}]^0 \right)$$

- The inverse mean map is given by

$$\begin{aligned} \psi_{\underline{s}}^{(M)}(m) &= \sum_{k=1}^{M-1} s_k (m_{\{k:k+1\}}^{-1})^0 + \sum_{k=M+1}^n s_k (m_{\{k-1:k\}}^{-1})^0 & (1) \\ &\quad - \sum_{k=2}^{M-1} s_{k-1} (m_{\{k,k\}}^{-1})^0 - (s_{M-1} - s_M + s_{M+1}) (m_{\{M,M\}}^{-1})^0 \\ &\quad - \sum_{k=M+1}^{n-1} s_{k+1} (m_{\{k,k\}}^{-1})^0. \end{aligned}$$

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SOME NOTATIONS

Let $m \in Q_G$; $\hat{m} \in S_n^+$ is the unique symmetric positive definite matrix verifying $\pi(\hat{m}) = m$, $\hat{m}^{-1} \in P_G$.

For all $m \in Q_G$, let $M_A = [(\hat{m}_A^{-1})^{-1}]^0$.

For all $A \in S_d$, let $\rho(A)u = AuA$ and write $\mathbb{P} = \pi \circ \rho$.

THEOREM 1

The variance function of the exponential family of Wishart distributions on Q_{A_n} is given by

$$\begin{aligned}
 V(m) = & \sum_{i=1}^{M-1} (s_i - s_{i+1}) \mathbb{P} \left(\sum_{j=1}^{i-1} \left(\frac{1}{s_j} - \frac{1}{s_{j+1}} \right) M_{1:j} + \frac{1}{s_i} M_{1:i} \right) \\
 & + s_M \mathbb{P} \left(\sum_{j=1}^{n-1} \left(\frac{1}{s_j} - \frac{1}{s_{j+1}} \right) M_{1:j} + \frac{1}{s_n} \hat{m} \right) \quad (2) \\
 & + \sum_{i=M+1}^n (s_i - s_{i-1}) \mathbb{P} \left(\frac{1}{s_i} M_{i:n} + \sum_{j=i+1}^n \left(\frac{1}{s_j} - \frac{1}{s_{j-1}} \right) M_{j:n} \right).
 \end{aligned}$$

THEOREM 2

The variance function of the Wishart exponential family $\gamma_{\underline{s}, y}^{(M)}$ is

$$\begin{aligned}
 V(m) = & \left[\frac{1}{s_1} + \frac{1}{s_n} - \frac{1}{s_M} \right] \mathbb{P}(\hat{m}) + \sum_{k=1}^{M-1} \left[\frac{1}{s_{k+1}} - \frac{1}{s_k} \right] \mathbb{P}(\hat{m} - M_{1:k}) \\
 & + \sum_{k=M+1}^n \left[\frac{1}{s_{k-1}} - \frac{1}{s_k} \right] \mathbb{P}(\hat{m} - M_{k:n}).
 \end{aligned}$$

IDEA OF PROOFS

The proof of the first result is based on writing

$V(m) = v(\psi_{\underline{s}}^{(M)}(m))$, where

$$v(y) = \sum_{i=1}^{M-1} (s_i - s_{i+1}) \mathbb{P} \left[((y_{\{1:i\}})^{-1})^0 \right] + s_M \mathbb{P}(y^{-1}) \\ + \sum_{i=M+1}^n (s_i - s_{i-1}) \mathbb{P} \left[((y_{\{i:n\}})^{-1})^0 \right].$$

Then we use a Choleski decomposition of $y = \psi_{\underline{s}}^{(M)}(m) = TT^T$, where the “triangular” matrix T has the same pattern of zeros as y .

The second result is proved by induction.